Section 5: Recovering Risk Types and (Risk) Preferences

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- Simple motivation: learn about preferences and types from choices and events
- Policy motivation: how bad is adverse selection and what should we do about it?

- 1. Infer (distribution of) risk types from risk realizations
- 2. Infer (distribution of) risk preferences with the above + (distribution of) choices
- 3. Run counterfactuals using (joint distribution of) risk types + preferences

- Review Cohen and Einav (2007) ECMA on car insurance choice
- Gain comfort with the idea of a model delivering

 $choice_i = f(risk type_i, risk preference_i)$

- Gain comfort with behavioral and functional form assumptions to recover model parameters from (imperfect) data
- Point out tips for digesting structural papers as "structural signposts"

Picture of Big Picture Intuition

Notation: consumer *i*, choice *j*, prices p_j , indirect utilities μ_{ij} , choice regions A_{ij}



Figure 1: Choice regions for goods 0, 1, and 2.

Source: Berry and Haile (2021) WP

Institutional Details

Model

Identification of Risk Preferences

- Unobserved heterogeneity: risk type and risk aversion
- Realized risks: accidents
- **Observed choices**: trading off premium (always pay) vs. deductible (pay only after accident) in menu of contracts

Unlike EFS, Cohen and Einav Have Price Variation!

$$d_{it} = min\{.5p_{it}, cap_t\}$$



FIGURE 1. VARIATION IN THE DEDUCTIBLE CAP OVER TIME

Firm says this was experimentation. Can avoid some assumptions w/ random variation...

- Annuity guarantee: Permanent decision at time of annuity purchase
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 ⇒ Consider "instantaneous" contract to isolate "static" demand
 - 1. **Tractable**: get simple closed form expressions for choice_i = $f(\text{risk type}_i, \text{risk preference}_i)$
 - 2. Connected to research question interested in risk preferences (not time preferences)
 - 3. Realistic: observe many cancellations in data

Institutional Details

Model

Identification of Risk Preferences

- Fully specify preferences to get $choice_i = f(risk type_i, risk preference_i)$
 - Identification problem: Choices driven by two unobserved dimensions
 - Identification solution: Make an assumption so one dimension is identified by something other than choice

- 1. Infer the distribution of risk types using observed risk realizations
 - What is the key assumption?
- 2. Given that, infer distribution of risk preferences from observed contract choices

- 1. Moral hazard
- 2. Non-random attrition due to early cancellation
- 3. Unreported accidents
- 4. Two dimensions of unobserved risk: frequency and size

(More or less) assume away!

- 1. Moral hazard
 - \rightarrow assume away!
- 2. Non-random attrition due to early cancellation
 - \rightarrow assume constant arrival rate and focus on per unit of time
- 3. Unreported accidents
 - \rightarrow assume threshold above which everything is reported
- 4. Two dimensions of unobserved risk: frequency and size
 - \rightarrow size is entirely idiosyncratic

- Main idea: infer riskiness of observable groups based on their risk realizations
- Main assumption: no moral hazard
- Additional implementation challenges: data censoring

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 - In our sample, the number of accidents is a noisy measure of accident probability
 - Fundamental tension between needing a very long time series and allowing age effects
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- What type of data would we need to infer whether a given person's **risk type changes** under different contracts?
 - Need to additionally see people under multiple contracts
 - Ideally contracts would be randomly assigned

1. Model accidents as Poisson process w/ parameter λ

2. Parametrize based on observables: $ln(\lambda_i) = x'_i\beta + \varepsilon_i$, where $\varepsilon_i \sim N(0, \sigma_{\lambda}^2)$

- 1. Model accidents as Poisson process w/ parameter λ
 - \blacksquare λ_i captures *i*'s "risk type"
 - Structural signpost # 1: Familiarize yourself with "go-to" distributions for contexts
- 2. Parametrize based on observables: $ln(\lambda_i) = x'_i\beta + \varepsilon_i$, where $\varepsilon_i \sim N(0, \sigma_{\lambda}^2)$
 - We don't get to observe λ_i , but we get its distribution from many people who look like i
 - Structural signpost # 2: N() is computationally convenient and often a decent descriptor of the population characteristics
 - Structural signpost # 3: log() accommodates parameters with sign restrictions
 - Structural signpost # 4: Incorporate heterogeneity based on observables

Claims	Low	Regular	High	Very high	Total	Share
0	11,929 (0.193)	49,281 (0.796)	412 (0.007)	299 (0.005)	61,921 (1.00)	0.8034
1	3,124 (0.239)	9,867 (0.755)	47 (0.004)	35 (0.003)	13,073 (1.00)	0.1696
2	565 (0.308)	1,261 (0.688)	4 (0.002)	2(0.001)	1,832 (1.00)	0.0238
3	71 (0.314)	154 (0.681)	1 (0.004)	0 (0.000)	226 (1.00)	0.0029
4	6 (0.353)	11 (0.647)	0 (0.000)	0 (0.000)	17 (1.00)	0.0002
5	1 (0.500)	1 (0.500)	0 (0.000)	0 (0.000)	2 (1.00)	0.00003

TABLE 2B—SUMMARY STATISTICS—CONTRACT CHOICES AND REALIZATIONS

- Eyeball the positive correlation test
- Structural signpost # 5: Look for summary stats that drive the model (n.b. see Andrews, Gentzkow, Shapiro (2020) ECMA for a formal treatment)

- Assume no moral hazard so that realizations reveal type
- Observe only one realization \rightarrow parametrize type based on observables
- Parameters: β, σ_{λ}

- **Main idea**: write down a model choice_i = $f(risk type_i, risk preference_i)$
- **Main assumptions**: choices are driven by inferred risk information and reveal underlying preferences
- Additional implementation challenges: discrete choices yield set identification rather than point identification

- Suppose we knew λ_i and choices are continuous
 - \Rightarrow can invert choice_i = $f(\text{risk type}_i, \text{risk preference}_i)$ to get exact risk preferences
- Discrete choices
 - \Rightarrow get bounds on risk preferences
- Observe λ_i 's distribution rather than exact value
 - \Rightarrow get distribution of risk preferences

- Consider utility for coverage length t and ake $\lim t \to 0$ for "instantaneous contract"
- Derive indiff. condition btw contracts in terms of risk type/aversion

Inferring Risk Preferences: "Instantaneous Contract"

Poisson:

$$P(k \text{ accidents over time } t) = \frac{(\lambda t)^k \exp(-\lambda t)}{k!}$$

• k > 1 terms vanish as $t \to 0$

• EU from contract price *p* and deductible *d* over small interval *t*:

$$v(p,d) \approx \underbrace{(1-\lambda t)}_{P(\text{no accident})} u(w-pt) + \underbrace{(\lambda t)}_{P(1 \text{ accident})} u(w-pt-d)$$

Inferring Risk Preferences: Indifference Condition

- Consider a high vs. low-deductible contract
 - Notation check on deductibles: $d^L < d^H \Rightarrow p^L > p^H$
- Indifference condition on contracts for the marginal type:

$$\boldsymbol{v}(\boldsymbol{p}^L,\boldsymbol{d}^L)=\boldsymbol{v}(\boldsymbol{p}^H,\boldsymbol{d}^H)$$

• Solve λ and take $t \to 0$ (see next slide for details):

$$\lambda = \frac{(p^L - p^H)u'(w)}{u(w - d^L) - u(w - d^H)}$$

• Rearrange to see MB vs. MC of low-deductible plan

Math Asides on Solving for λ

- Substitute for v()'s and divide through by t
- p^H, p^L disappear in terms with t only in $u(\cdot)$ since $\lim_{t\to 0} p^H t = \lim_{t\to 0} p^L t = 0$
- Need to express terms with *t* as derivative:

$$\begin{aligned} &\frac{1}{t} \left[u(w - p^{H}t) - u(w - p^{L}t) \right] \\ &= \frac{1}{t} \left[\left(u(w - p^{H}t) - u(w) \right) - \left(u(w - p^{L}t) - u(w) \right) \right] \\ &= p^{H} \frac{u(w - p^{H}t) - u(w)}{p^{H}t} - p^{L} \frac{u(w - p^{L}t) - u(w)}{p^{L}t} \\ &= (p^{L} - p^{H})u'(w) \end{aligned}$$

- Previously backed out (the distribution of) risk type λ
- Derived locus of risk type/aversion indifferent btw contracts:

$$\lambda = \frac{(p^L - p^H)u'(w)}{u(w - d^L) - u(w - d^H)}$$

- What can we do with an expression containing $u(\cdot)$, $u'(\cdot)$, and w?
 - Recall Baily-Chetty formula that mapped unobservable u'(c) gap into observables

Inferring Risk Preferences: From hopeless $u(\cdot)$ to hopeful r

• Take 2^{nd} order Taylor expansion of u around w in previous expression

$$u(w - d^L) \approx u(w) - u'(w)d^L + \frac{1}{2}u''(w)[d^L]^2$$

- When is this exact?
- Structural signpost # 6: Don't miss forest through trees. Goal is connecting model to data.
- Recall coefficient of **absolute** risk aversion $r(w) = -\frac{u''(w)}{u'(w)}$
 - EFS assumed CRRA because choice was over fraction of wealth
 - Cohen and Einav assume CARA because choice is over dollar amount
- Do algebra in the privacy of your own home:

$$r^*(\lambda) = \frac{\frac{p^L - p^H}{\lambda(d^H - d^L)} - 1}{\frac{1}{2}(d^H - d^L)}$$

Recap: What is this telling us?

Institutional Details

Model

Identification of Risk Preferences

- We made an behavioral assumption (no MH) and functional form assumption (Poisson parameters distributed lognormal) to *identify* risk types
- We derived a model to get $choice_i = f(risk type_i, risk preference_i)$
- I'll have a brief digression about model *identification*
- Then we will discuss assumptions that identify risk preferences

What does model identification mean?

What does model identification mean?

- A model is (set) identified if different (sets of) values of the parameters imply different distributions of observable data (Matzkin 2013 ARE)
- Identification is a binary property of a(n economic or econometric) model

- 1. Econometric model: Additively linear in age, calendar time, and cohort
 - Age = calendar time cohort
 - Different individual parameters consistent with same distribution of observable data ⇒ not identified (Ameriks and Zeldes 2004)
- 2. Economic model: Equilibrium relationship between supply and demand
 - Supply (demand) shifters identify the demand (supply) curve

Identifying Risk Preferences: Illustration

Indifference condition from before: $r^*(\lambda) = 2 \left[\frac{1}{\lambda} \frac{\Delta p}{(\Delta d)^2} - \frac{1}{\Delta d} \right]$



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1. What does variation in Δp , Δd , and $\Delta p/\Delta d$ do? (www.desmos.com/calculator)

2. What variation and outcomes do we observe in the data?

- Given λ , choice identifies a **set** of possible r
- Random variation in prices and menus delivers many such sets

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- Random variation in prices and menus delivers many such sets
- Sufficient variation identifies r without additional functional form assumptions
 - Intuition: Find out a number by repeatedly asking if it's > x for many different x

- 1. Cohen and Einav: Don't actually observe infinite price variation
- 2. EFS: Don't observe any price variation, so can get only identified sets without further assumptions
- 3. Structural signpost #7: Keep track of what assumptions are required by setting vs. want of point identification/lack of infinite data vs. lack of random variation

Parametric Identification: Setup

• Recall parametric assumption on risk type:

$$\ln \lambda_i = x_i'\beta + \varepsilon_i$$

where $\varepsilon \sim N(0, \sigma_{\lambda})$

• Additionally make parametric assumption on risk aversion:

$$\ln r_i = x_i'\gamma + v_i$$

where $v_i \sim N(0, \sigma_r)$

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- Note: We actually care about $\mu_{\lambda} \equiv E_i[\lambda_i]$ and $\mu_r \equiv E_i[r_i]$ rather than β and γ
- Allow $Cov(\varepsilon_i, v_i) = \rho \sigma_\lambda \sigma_r$ since motivation is joint distribution of unobservables

- Model parameters: risk types $\mu_{\lambda}, \sigma_{\lambda}$; risk preferences μ_r, σ_r ; and correlation ρ
- Rough intuition: Need 5 relevant moments to identify 5 parameters

- Fractions with k claim together identify $\mu_{\lambda},\sigma_{\lambda}$
- Fractions choosing low deductible among those with 0,1, and 2 claims identify the remaining preference and correlation parameters
- Define $\phi_k \equiv$ fraction who chose the low deductible plan among those who realized k claims for $k \in \{0, 1, 2\}$

Parametric Identification: Graphical Intuition



Parametric Identification: Graphical Intuition



Sequential thought experiments:

- 0. We have $\mu_{\lambda}, \sigma_{\lambda}$
- 1. Suppose r constant (i.e.
 - $\sigma_r = \rho = 0) \rightarrow \text{bar height}$
- 2. Now suppose $\sigma_r > 0$ but $\rho = 0$ \rightarrow bar slope
- 3. Now suppose $\sigma_r, \rho > 0 \rightarrow$ bar convexity

- Collect parameters: $\Theta = \{\beta, \sigma_{\lambda}, \gamma, \sigma_{r}, \rho\}$
- Write down likelihood of observed choices given a candidate Θ :

 $L(claims_i, choice_i | \Theta) = Pr(claims_i, choice_i | \lambda_i, r_i) Pr(\lambda_i, r_i | \Theta)$

- Maximize likelihood?
 - Turns out this is computationally hard
 - (Evaluating likelihood once requires integrating over both λ_i and r_i for every i)
 - Gibbs Markov Chain Monte Carlo to the rescue!

Gibbs Sampling General Intuition



Both methods involve the data disciplining parameter estimation

Simulation methods: For when you/your computer is too dumb to evaluate something

- MLE \approx frequentist, Gibbs MCMC \approx Bayesian
- Gibbs procedure:
 - 1. Take draws of all parameters from priors
 - 2. Draw a single parameter from a posterior (given observables and other drawn parameters)
 - 3. Do previous step with a different parameter
 - 4. Continue iterating over all parameters many times

Crazy result: The sequence of posterior draws converges to the joint distribution¹ **Upshot**: After a bunch of iterations, averaging over many subsequent draws delivers (mean) parameter estimate

- Given (parameters governing distribution of) r_i , observed choices tell you which (parameters governing distribution of) λ_i are likely
- Given those (parameters governing distribution of) λ_i , (parameters governing distribution of) r_i are likely
- Applying the discipline of observed choices many times eventually delivers parameters "close to the truth"

- Separately identify multidimensional unobserved types: risk type and preference
 - "Looking under the hood" of the positive correlation test
 - Common themes in "structural" insurance papers:
 - 1. Use ex post realizations to infer ex ante risk
 - 2. Make assumption on contract choice process
- Requires assumptions to see what objects in data can be mapped to unobservables
- Also requires a lot of structure. In my view, the paper transparently:
 - 1. Argues why assumptions and structure are necessary
 - 2. Shows where identification comes from
 - 3. Focuses on an interesting question without getting distracted